



Vitamin D Deficiency and Comorbidities as Risk Factors of COVID-19 Infection: A Systematic Review and Meta-analysis

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Objectives: Extensive evidence links low vitamin D status and comorbidities with coronavirus disease 2019 (COVID-19) outcomes, but the results of published studies are contradictory. Therefore, we investigated the association of lower levels of vitamin D and comorbidities with the risk of COVID-19 infection.

Methods: We searched MEDLINE (via PubMed), Cochrane Central Register of Controlled Trials, and ClinicalTrials.gov for articles published until August 20, 2021. Sixteen eligible studies were identified (386 631 patients, of whom 181 114 were male). We included observational cohort and case-control studies that evaluated serum levels of vitamin D in COVID-19-positive and COVID-19-negative patients. Mean differences (MDs) with 95% confidence intervals (CIs) were calculated.

Results: Significantly lower vitamin D levels were found in COVID-19-positive patients (MD, -1.70; 95% CI, -2.74 to -0.66; $p=0.001$), but with variation by study design (case-control: -4.04; 95% CI, -5.98 to -2.10; $p<0.001$; cohort: -0.39; 95% CI, -1.62 to 0.84; $p=0.538$). This relationship was more prominent in female patients (MD, -2.18; 95% CI, -4.08 to -0.28; $p=0.024$) than in male patients (MD, -1.74; 95% CI, -3.79 to 0.31; $p=0.096$). Male patients showed higher odds of having low vitamin D levels (odds ratio [OR], 2.09; 95% CI, 1.38 to 3.17; $p<0.001$) than female patients (OR, 1.17; 95% CI, 0.74 to 1.86; $p=0.477$). Comorbidities showed inconsistent, but generally non-significant, associations with COVID-19 infection.

Conclusions: Low serum vitamin-D levels were significantly associated with the risk of COVID-19 infection. This relationship was stronger in female than in male COVID-19 patients. Limited evidence was found for the relationships between comorbidities and COVID-19 infection, warranting large population-based studies to clarify these associations.

Key words: Vitamin D, COVID-19, Comorbidities, Risk factors, Meta-analysis

INTRODUCTION

The world is in the grasp of the coronavirus disease 2019 (COVID-19) pandemic, which began in late 2019 in Wuhan,

Received: December 18, 2021 Accepted: May 9, 2022

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China. The infectious agent responsible for COVID-19 was originally called 2019-nCoV; the disease was renamed COVID-19, and its causative agent was renamed severe acute respiratory syndrome coronavirus 2 (SARS-CoV-2), by the World Health Organization on February 11, 2020. Previous coronavirus epidemics included that of severe acute respiratory syndrome SARS-CoV, initiated in China in 2002, and that of Middle East respiratory syndrome (MERS)-CoV, first recorded in 2012. These epidemics all started with the spread of the infection between animals and humans. The primary cause of death is usually due to extreme atypical pneumonia [1]. To date, there is no established curative therapy for this virus, and preven-

tion remains the best strategy for combating the COVID-19 pandemic.

Vitamin D supplementation is considered to be a preventive strategy, as indicated by certain observational studies [2-4]. Insufficiency of vitamin D is a public health problem affecting over a billion people across all life stages worldwide [5]. In the past decade, several studies have demonstrated a potential link between vitamin D deficiency and various diseases, including systemic infections [6-8]. Various clinical studies have reported associations of low serum vitamin D levels with acute respiratory tract infections, including epidemic influenza [9-11]. Vitamin D, a steroid hormone, whose biosynthesis begins with solar ultraviolet radiation in bare skin exposed to strong sunlight, and shows multidimensional effects beyond calcium and bone metabolism. Vitamin D improves mucosal defenses by secreting antiviral peptides [12,13]. Vitamin D receptors are highly expressed in B and T lymphocytes, suggesting that they may play a role in modulating the innate and adaptive immune responses [14].

Vitamin D levels can be affected by many factors such as sun exposure, genetics, supplementation, and comorbidities. Vitamin D levels decrease during winter, and low vitamin D levels are associated with an increased risk of acute respiratory tract infections during winter [15] mitigated by vitamin D supplementation. Several well-defined conditions have been recognized as risk factors for a worsening disease course and poor COVID-19 outcomes. The most important risk factors that may result in severe COVID-19 symptomatology are diabetes mellitus, hypertension, age over 65 years, obesity, and immunosuppressive therapy. By diverse mechanisms, these circumstances are hypothesized to alter host responses to infection, enhancing and ramping up harmful pathophysiological processes. Initial viral immune evasion and the ensuing hyperinflammatory response resulting from excessive and undirected immune activation are crucial components in COVID-19 pathogenesis [16,17]. People older than 60 years of age with hypertension, diabetes, and respiratory, cardiovascular, cerebrovascular, liver, kidney and gastrointestinal disorders are more vulnerable to COVID-19 infection and experience higher mortality. Because of the limited number of patients, the involvement of malignant conditions is under debate [18].

Extensive evidence has recently linked low vitamin D status with COVID-19 outcomes; however, these results are contradictory: 2 retrospective studies reported independent associations between low pre-pandemic vitamin D levels and the

subsequent incidence and severity of COVID-19 [19,20], while an analogous study in the United Kingdom did not support the potential link between vitamin D concentration and the risk of severe COVID-19 infection and mortality [21]. A recent meta-analysis integrating data from 8 observational studies reported an increased risk of community-acquired pneumonia in patients with a serum vitamin D concentration <20 ng/mL [22]. Some recent reviews hypothesized that vitamin D insufficiency may compromise respiratory immune function, increasing the risk of COVID-19 [1,23]. Some retrospective studies have also observed possible correlations of vitamin D levels with COVID-19 outcomes [21,24-29]. Several studies have investigated the effect of vitamin D supplementation on COVID-19 outcomes [2] and evaluated the risk of developing COVID-19 infection in vitamin D-deficient patients and those with normal vitamin D levels [30,31], indicating the possible existence of a link between vitamin D levels and the risk of COVID-19 infection. However, the available data continue to be an area of uncertainty and an ongoing focus of attention [30].

Low vitamin D levels are associated with comorbidities that are known to affect COVID-19 outcomes. Further investigations should focus on patients with low vitamin D levels with or without comorbidities and supplementation trials to investigate the effects of vitamin D on the immune response to COVID-19. Therefore, in the present systematic review, we aimed to determine the association of lower levels of vitamin D and comorbidities with the risk of COVID-19 infection. We hypothesized that lower vitamin D levels in the blood would be linked to a higher risk of COVID-19 infection.

METHODS

The Preferred Reporting Items for Systematic Review and Meta-Analysis (PRISMA) guidelines for systematic reviews [32] and MOOSE (Meta-analysis of Observational Studies in Epidemiology) guidelines [33] were followed for designing, conducting, and reporting this systematic literature review. The protocol of this systematic review was registered in PROSPERO (registration ID: CRD42020205150).

Data Sources and Searches

Three independent reviewers conducted the search for studies in the MEDLINE (via PubMed), Cochrane Central Register of Controlled Trials and ClinicalTrials.gov until August 20, 2021, using merged Medical Subject Headings (MeSH) and non-MeSH

terms as follows: “vitamin D” OR “25-hydroxy vitamin D” AND “COVID-19” OR “2019 novel coronavirus” OR “SARS CoV-2”. We also searched the gray literature using Google Scholar and manually checked the reference lists of eligible articles.

Inclusion and Exclusion Criteria

Studies reporting the levels of vitamin D and comorbid conditions in COVID-19-positive and COVID-19-negative patients were included. We excluded reviews, editorials, opinions, case reports, case-series, perspectives, letters, protocols, and studies not reporting the required data. The first author (PM) searched data and screened articles for eligibility. The senior author (RP) double-checked all the included articles, and any dispute was resolved by the third author (RB).

Quality Assessment

Two reviewers (PM and RP) assessed the quality of data in the included studies using the National Institute of Health (NIH) quality assessment tools developed by the National Heart, Lung, and Blood Institute (NHLBI) [34]. The NIH tool was preferred because it is comprehensive and widely accepted for an exhaustive assessment of data quality. The tools were designed to assist reviewers in focusing on concepts that are key for critical appraisal of the internal validity of a study. The tools were not designed to provide a list of factors comprising a numeric score. The tools were specific to individual types of included study designs and are described in more detail below. The tools included items for evaluating potential flaws in study methods or implementation, including sources of bias (e.g., patient selection, performance, attrition, and detection), confounding, study power, the strength of causality in the association between interventions and outcomes, and other factors. The quality reviewers could select “yes,” “no,” or “cannot determine/not reported/not applicable” in response to each item on the tool. For each item where “no” was selected, reviewers were instructed to consider the potential risk of bias that could be introduced by that flaw in the study design or implementation. Responses of “cannot determine” and “not reported” were also noted as representing potential flaws. Each of the quality assessment tools had a detailed guidance document, which was also developed by the methodology team and NHLBI.

Data Extraction

Data were entered into a standardized data extraction table (Excel) and independently checked by a second reviewer (RP)

for accuracy. The following raw data were extracted: name of the first author, year of publication, country of origin, study design, age, sex, identified comorbidities, reported levels of vitamin D (in the form of mean and standard deviation), patients’ socioeconomic status in terms of the Townsend deprivation quintile [21], and the number of patients in the COVID-19-positive and COVID-19 negative groups. The included studies designated reverse-transcription polymerase chain reaction (RT-PCR)–confirmed patients as COVID-19 positive and RT-PCR negative patients as COVID-19 negative groups. The patients were categorized as vitamin D-deficient or having low vitamin D levels if their vitamin D levels were <10 ng/mL (<25 nmol/L) [21,35], <20 ng/mL (<50 nmol/L) [5,26,36–43], <12 ng/mL (<30 nmol/L) [44,45] or <30 ng/mL (<75 nmol/L) [46,47]. Comorbidities described using words or phrases such as “hypertension” and “high blood pressure,” were classified as using the term “hypertension.” Comorbidities described as “chronic obstructive pulmonary disease,” “COPD,” or “chronic lung diseases” were referred to as “chronic lung disorder” in our study.

Data Synthesis

We performed an exploratory meta-analysis to understand the magnitude and direction of the effect estimate. Continuous outcomes are presented using weighted mean differences (MDs) and 95% confidence intervals (CIs). Odds ratios (ORs) were calculated and presented with respective 95% CIs for binary outcomes. The Mantel-Haenszel method for binary outcomes and the inverse-variance method for continuous outcomes were used to calculate 95% CIs. A random-effects model with the DerSimonian-Laird method was used to pool effect estimates, as substantial methodological heterogeneity was observed when pooling effect estimates [48]. Heterogeneity between studies was assessed using the chi-square-based Cochran Q statistic (with $p < 0.1$ considered as indicating the presence of heterogeneity) and the I^2 statistic (with >50% representing moderate heterogeneity) [48]. Publication bias was assessed only for the primary outcome by a visual inspection of a funnel plot, as the requirement for the minimum number of studies (≥ 10 studies) was satisfied. The Egger regression test was applied to assess small-study effect (with $p < 0.1$ considered as indicating the presence of the small-study effect). We also used the Duval and Tweedie trim-and-fill method to estimate what the summary effect size would be if there was no publication bias [49]. A subgroup analysis was conducted if vitamin D deficiency was reported in proportions. We also ana-

lyzed the corresponding data if the mean vitamin D levels were reported by sex. We also calculated 95% prediction intervals where 3 or more studies were available; these intervals represent the direction and range of an effect estimate in a new study [50]. All statistical analyses were conducted using Stata version 17.0 (StataCorp., College Station, TX, USA), and a *p*-value less than 0.05 was considered a statistically significant result.

Ethics Statement

As the present study was a meta-analysis, the data was extracted from the published articles. Therefore, institutional review board approval was not required.

RESULTS

Study Characteristics

The systematic search yielded a total of 1723 publications. After removing duplicates, 611 articles were found to be potential publications for screening. Studies were published between January 1, 2020 and August 20, 2021. After applying

predefined inclusion and exclusion criteria, a total of 16 studies were included in the qualitative and quantitative analyses (Figure 1). The included studies were cohort studies [20,21, 25,26,39,40,44-46,51] and case-control studies [35-38,41,52]. The included 16 studies enrolled a total of 386 631 patients from 4 different countries, including 181 114 male patients and 205 517 female patients. The background characteristics of the included studies are presented in Table 1.

Eight of the included articles [21,35-38,41,46,51] presented specific comorbidity data, in which the 4 most prevalent comorbidities were diabetes, cardiovascular disease, respiratory diseases and hypertension. The most common cardiovascular diseases were found to be coronary artery disease and hypercholesterolemia, while chronic lung disorder was the most prevalent respiratory disease. We also included the Townsend deprivation index in our analysis as a comorbidity.

Quality Assessment

We assessed the quality of data in the included studies using the NIH quality assessment tools, and presented the results in

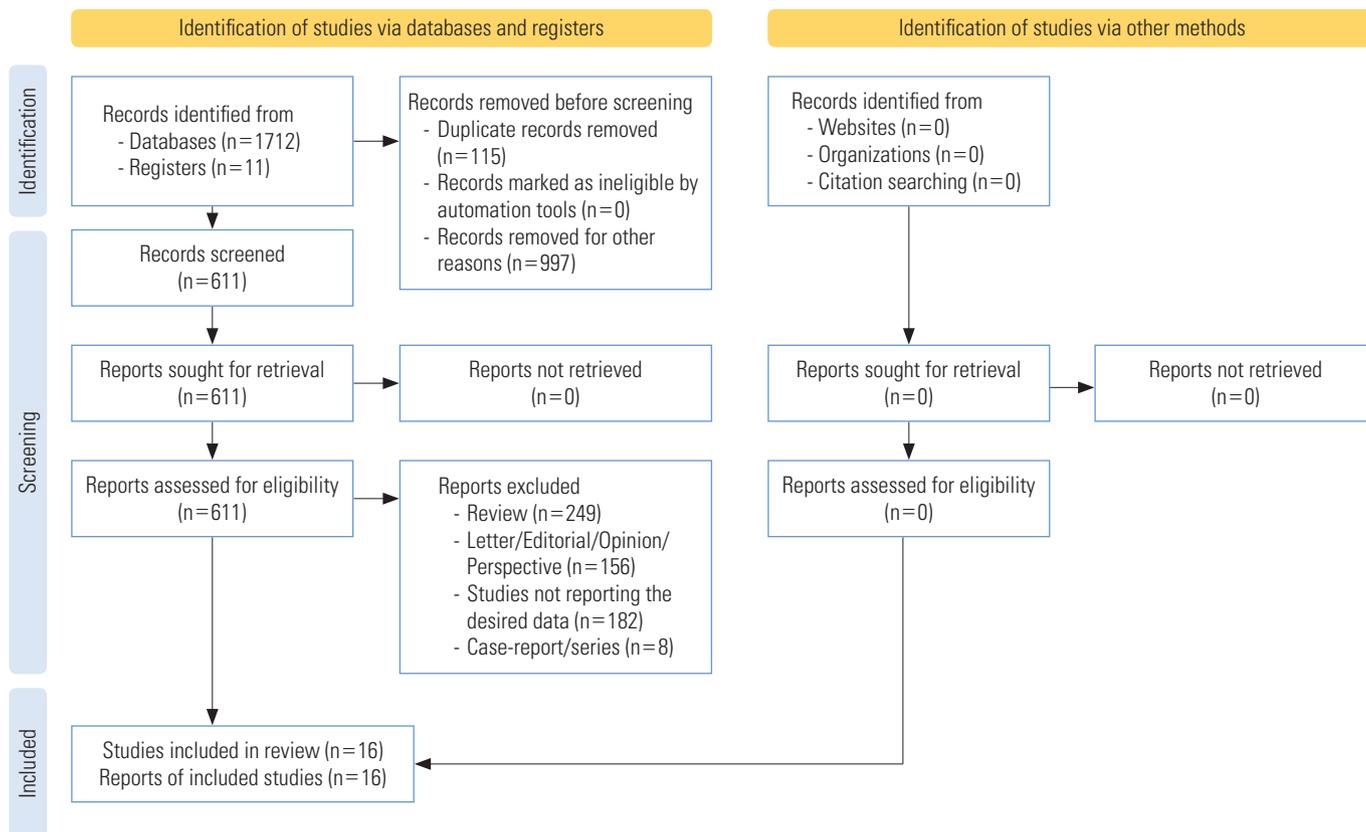


Figure 1. PRISMA (Preferred Reporting Items for Systematic Review and Meta-Analysis) flow diagram of the number of studies screened and included in the systematic review and meta-analysis.

Table 1. Characteristics of study participants from the included studies

Study	Study design	Country	Groups	Sample size	Age, mean ± SD or median (IQR)/range y	Sex, male (%)	Definition of vitamin D status	Levels of vitamin D, mean ± SD or median (IQR)/range	Comorbidities	Quality Index
D'Avolio et al., 2020 [20]	Retrospective cohort	Switzerland	COVID-19 positive COVID-19 negative	27 80	74 (65-81) 73 (61-82)	19 (70.4) 39 (48.8)	NR	11.1 (8.2-21.0) 24.6 (8.9-30.5)	NA	Fair
Darling et al., 2020 [25]	Prospective cohort	UK	COVID-19 positive COVID-19 negative	580 723	57.5 (8.7) 57.9 (8.7)	336 (57.5) 377 (52.1)	NR	43.3 (32.1) 44.1 (31.2)	NA	Fair
Raisi-Estabragh et al., 2020 [51]	Prospective cohort	UK	COVID-19 positive COVID-19 negative	1326 3184	68.11 ± 9.23 68.91 ± 8.72	696 (52.5) 1,505 (47.3)	NR	33.88 ± 27.01 35.45 ± 26.78	Diabetes, hypertension, high cholesterol	Good
Hastie et al., 2020 [21]	Retrospective cohort	UK	COVID-19 positive COVID-19 negative	449 348 149	49 (40-58) 49 (38-57)	265 (59.0) 168 391 (48.4)	Deficiency (<25 nmol/L), Insufficiency (<50 nmol/L)	28.7 (10.0-43.8) 32.7 (10.0-47.2)	Diabetes	Good
Merzon et al., 2020 [46]	Retrospective cohort	Israel	COVID-19 positive COVID-19 negative	782 7025	68.11 ± 9.23 68.91 ± 8.72	385 (49.2) 397 (50.8)	Low (<30 ng/mL)	19.00 (18.41-19.59) 20.55 (20.32-20.78)	Diabetes, hypertension, high cholesterol, depression, schizophrenia, dementia, cardiovascular disease, coronary artery disease, chronic lung disorder	Good
De Smet et al., 2020 [26]	Retrospective cohort	Belgium	COVID-19 positive COVID-19 negative	186 2717	69 (52-80) 68 (49-82)	109 (58.6) 999 (36.8)	Deficiency (<20 ng/mL)	18.6 (12.6-25.3) 21.5 (13.9-20.8)	Cardiovascular disease, coronary artery disease,	Good
Hernández et al., 2021 [38]	Retrospective case-control	Spain	COVID-19 positive COVID-19 negative	197 197	61 (47.5-70) 61 (56-66)	123 (62.4) 123 (62.4)	Deficiency (<20 ng/mL)	13.8 ± 7.2 20.9 ± 7.4	Hypertension, diabetes, cardiovascular disease, chronic lung disease	Fair
Abdollahi et al., 2021 [35]	Case-control	Iran	COVID-19 positive	201	24 (19-29)	66	Deficient (<10 ng/mL), insufficient (10-30 ng/mL), sufficient (>30-100 ng/mL)	24 (19-29)	Hypertension, diabetes, chronic lung disease	Fair
Sulli et al., 2021 [41]	Case-control	Italy	COVID-19 positive COVID-19 negative	65 65	7.9 (15) 16.3 (19)	30 (46.2) 30 (46.2)	Sufficiency (>30 ng/mL) Insufficiency (between 20 and 30 ng/mL), Deficiency (between 10 and 20 ng/mL), and severe deficiency (<10 ng/mL)	26 (21-35) 7.9 (15) 16.3 (19)	Hypertension, diabetes, cardiovascular disease, coronary artery disease, chronic lung disease	Fair
Alguwaihes et al., 2021 [36]	Retrospective case-control	Saudi Arabia	COVID-19 positive COVID-19 negative	150 72	35.8 ± 1.5 42.5 ± 3.0	97 (64.7) 38 (52.8)	Deficiency (<50 nmol/L) and severe deficiency (<12.5 nmol/L)	35.8 ± 1.5 42.5 ± 3.0	Hypertension, diabetes, cardiovascular disease	Fair
Livingston et al., 2021 [45]	Retrospective cohort	UK	COVID-19 positive COVID-19 negative	47 57	38.9 ± 28.2 51.0 ± 31.4	20 (42.6) 19 (33.3)	Deficiency (<30 nmol/L)	38.9 ± 28.2 51.0 ± 31.4	NA	Fair
Gaudio et al., 2021 [44]	Retrospective cohort	Italy	COVID-19 positive COVID-19 negative	50 100	12.5 (2-42) 20.5 (5-46)	26 (52.0) 44 (44.0)	Deficiency <12 ng/mL (30 nmol/L) and insufficiency 12-20 ng/mL (50 nmol/L)	12.5 (2-42) 20.5 (5-46)	NA	Good

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Table 1. Continued from the previous page

Study	Study design	Country	Groups	Sample size	Age, mean ± SD or median (IQR)/range y	Sex, male (%)	Definition of vitamin D status	Levels of vitamin D, mean ± SD or median (IQR)/range	Comorbidities	Quality Index
Ferrari et al., 2021 [39]	Retrospective cohort	Italy	COVID-19 positive COVID-19 negative	188 1185	25.1 ± 13.2 26.7 ± 13.3	101 (53.7) 547 (46.1)	Deficiency (<20 ng/mL)	25.1 ± 13.2 26.7 ± 13.3	NA	Fair
Al-Daghri et al., 2021 [37]	Case-control	Saudi Arabia	COVID-19 positive COVID-19 negative	138 82	55.0 ± 28.8 61.8 ± 22.8	79 (57.2) 41 (50.0)	Deficiency (<50 nmol/L)	55.0 ± 28.8 61.8 ± 22.8	Hypertension, diabetes, high cholesterol, cardiovascular disease	Fair
Li et al., 2021 [40]	Cohort	USA	COVID-19 positive COVID-19 negative	900 17 248	25 (18-33) 27 (20-36)	252 (28.0) 5726 (33.2)	Levels below 20 or 30 ng/mL as low	25 (18-33) 27 (20-36)	NA	Good
Raesi et al., 2021 [52]	Case-control	Iran	COVID-19 positive COVID-19 negative	91 169	73.16 ± 23.59 76.02 ± 23.48	55 (60.4) 113 (66.9)	NR	73.16 ± 23.59 76.02 ± 23.48	Hypertension, diabetes, high cholesterol, cardiovascular disease, chronic lung disorder	Fair

COVID-19, coronavirus disease 2019; NA, not available; NR, not reported SD, standard deviation; IQR, interquartile range.

Table 1. The majority of the included studies (nearly 67%) were of acceptable quality. All the papers clearly stated the research question or objective, the study population was clearly specified and defined, and all the patients were selected from the same or similar populations. The detailed results of the quality assessment are provided in Supplemental Material 1.

Association Between Vitamin D Levels and COVID-19

The pooled estimate from 16 studies showed significantly lower serum levels of vitamin D in COVID-19-positive patients (MD, -1.70; 95% CI, -2.74 to -0.66; $p=0.001$) with a 95% prediction interval of -5.85 to 2.45 (Figure 2). When stratified by the study design, only the case-control design (MD, -4.04; 95% CI, -5.98 to -2.10; $p<0.001$) showed a significant difference between the COVID-19-positive and COVID-19-negative patients, unlike cohort studies (MD, -0.39; 95% CI -1.62 to 0.84; $p=0.538$). Substantial overall statistical heterogeneity ($I^2=94.2%$; Cochran Q test $p<0.001$) was observed between the studies. Four studies reported vitamin D deficiency by categorizing vitamin D levels. The pooled OR for vitamin D deficiency in COVID-19-positive patients was 1.61 (95% CI, 1.20 to 2.17; $I^2=56.3%$; $p=0.002$) compared to COVID negative patients, with a 95% prediction interval of 0.53 to 4.94.

Subgroup Analysis

Five studies reported vitamin D levels by sex (Figure 3). In a subgroup analysis based on sex, the difference in serum vitamin D levels was larger between female COVID-19 patients and female controls (MD, -2.18; 95% CI, -4.08 to -0.28; $p=0.024$; $I^2=37.5%$) than the corresponding difference between male patients and controls (MD, -1.74; 95% CI, -3.79 to 0.31; $p=0.096$; $I^2=67.0%$), with wide prediction intervals for both estimates. One study reported the proportion of vitamin D deficiency (vitamin D <20 ng/mL), and male patients showed higher odds of having low vitamin D levels (OR, 2.09; 95% CI 1.38 to 3.17; $p<0.001$) than female patients (OR, 1.17; 95% CI 0.74 to 1.86; $p=0.477$). When stratified by the study design, only a case-control design (MD, -4.04; 95% CI, -5.98 to -2.10; $p<0.001$) showed a significant difference between the COVID-19-positive and COVID-19-negative patients, unlike the cohort studies (MD, -0.39; 95% CI, -1.62 to 0.84; $p=0.538$), which had wider prediction intervals, as shown in Figure 2.

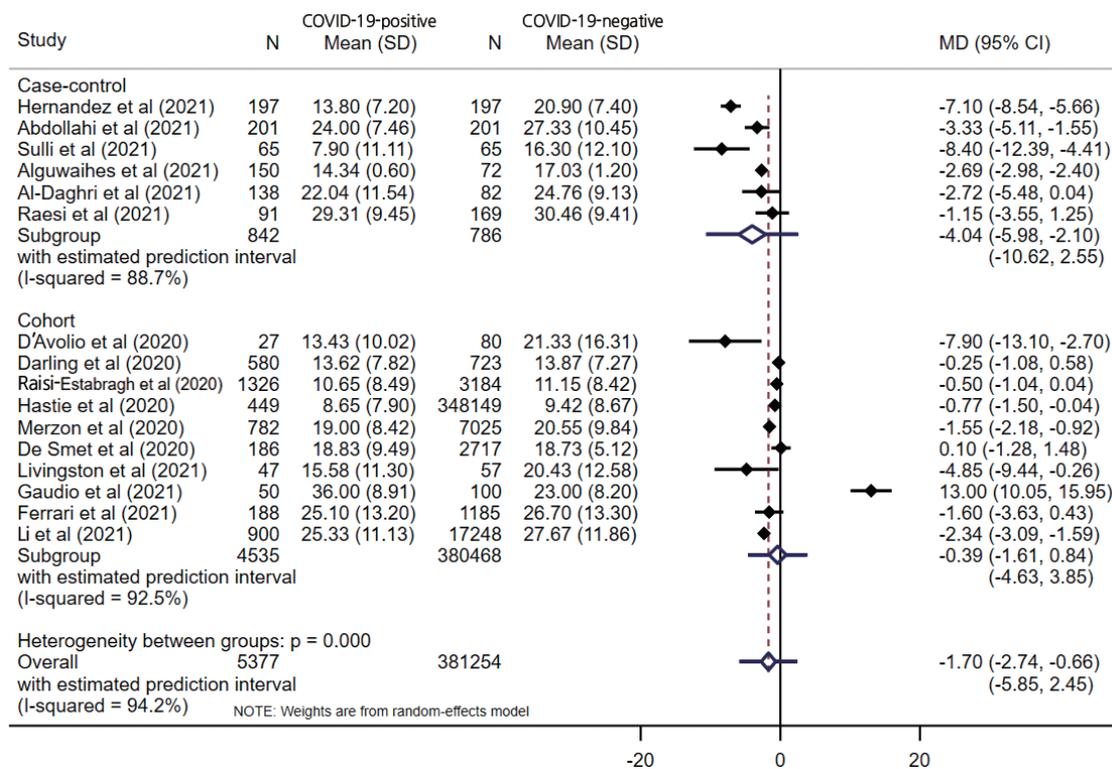


Figure 2. Forest plot showing a pooled analysis of vitamin D measurements in COVID-19 patients. COVID-19, coronavirus disease 2019; SD, standard deviation; MD, mean differences; CI, confidence interval.

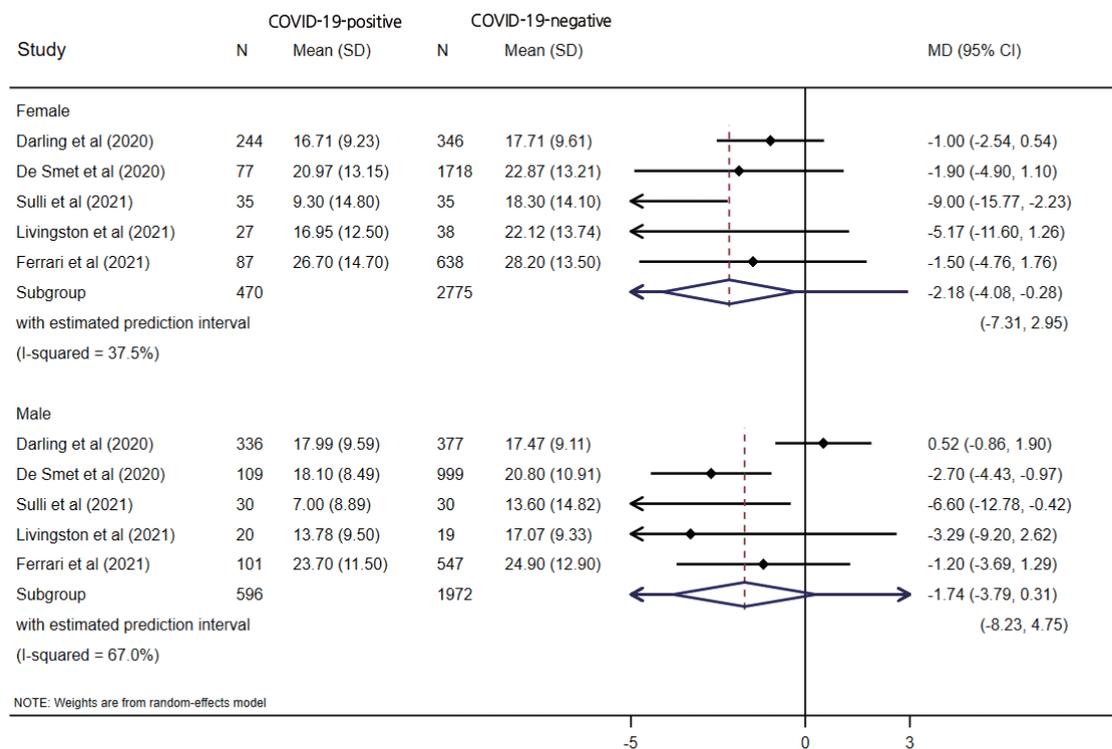


Figure 3. Forest plot showing a sex-specific pooled analysis of vitamin D measurements in COVID-19 patients. COVID-19, coronavirus disease 2019; SD, standard deviation; MD, mean differences; CI, confidence interval.

Table 2. Pooled summary for the association of comorbidities with COVID-19 infection

Comorbidity	No. of included studies/total studies	Group n/N		OR (95% CI)	p-value	95% prediction interval	
		COVID-19 positive	COVID-19 negative			UL	LL
Underweight	1/16	2/449	1759/348 149	0.88 (0.22, 3.54)	0.858	-	-
Overweight	2/16	272/650	148 610/348 350	1.58 (0.65, 3.86)	0.314	-	-
Obesity	1/16	158/449	82 770/348 149	1.74 (1.43, 2.11)	<0.001	-	-
Diabetes	8/16	1013/3308	331 477/358 939	1.14 (0.72, 1.82)	0.579	0.23	5.74
Hypertension	7/16	1045/2859	3615/10 826	0.99 (0.70, 1.43)	0.982	0.32	3.12
High cholesterol	2/16	510/2246	1851/10 209	0.95 (0.56, 1.63)	0.780	0.02	46.65
Cardiovascular disease	4/16	60/550	47/416	1.20 (0.39, 3.72)	0.754	0.01	159.25
Depression	1/16	73/782	141/7025	0.96 (0.56, 1.63)	0.866	-	-
Schizophrenia	1/16	15/782	427/7025	0.56 (0.37, 0.82)	0.003	-	-
Dementia	1/16	27/782	1172/7025	0.56 (0.43, 0.71)	<0.001	-	-
Coronary artery disease	2/16	77/847	946/7090	0.64 (0.46, 0.87)	0.005	-	-
Chronic lung disorder	4/16	272/1245	1920/7488	1.26 (0.99, 1.61)	0.055	0.63	2.54
Townsend deprivation quintile							
1	1/16	61/449	70 726/348 149	0.62 (0.47, 0.81)	<0.001	-	-
2	1/16	76/449	70 644/348 149	0.80 (0.62, 1.02)	0.075	-	-
3	1/16	64/449	70 270/348 149	0.65 (0.50, 0.85)	0.002	-	-
4	1/16	105/449	65 840/348 149	1.21 (0.97, 1.50)	0.092	-	-
5	1/16	143/449	65 840/348 149	2.00 (1.64, 2.44)	<0.001	-	-

COVID-19, coronavirus disease 2019; n, number of patients with the comorbidity; N, total patients in the group; OR, odds ratio; CI, confidence interval.

Comorbidities and Acquiring COVID-19 Infection

A pooled analysis of the data from the aforementioned studies showed that patients exposed to comorbidities such as obesity (OR, 1.74; 95% CI, 1.43 to 2.11; $p < 0.001$) and Townsend deprivation quintile 5 (OR, 2.00; 95% CI, 1.64 to 2.44; $p < 0.001$) had higher odds of acquiring COVID-19 infection. However, patients with schizophrenia (OR, 0.56; 95% CI, 0.37 to 0.82; $p = 0.003$), coronary artery disease (OR, 0.64; 95% CI, 0.46 to 0.87; $p = 0.005$), dementia (OR, 0.56; 95% CI, 0.43 to 0.71; $p < 0.001$), Townsend deprivation quintile 1 (OR, 0.62; 95% CI, 0.47 to 0.81; $p = 0.001$), Townsend deprivation quintile 3 (OR, 0.65; 95% CI, 0.50 to 0.85; $p = 0.002$) had lower odds of acquiring COVID-19 infection. Furthermore, other comorbidities such as diabetes, hypertensions, chronic lung disease, and cardiovascular disease showed no statistically significant associations with COVID-19 infection (Table 2).

Publication Bias

Publication bias was assessed for the primary outcome (i.e., overall vitamin D levels) by visual inspection of a funnel plot and the Egger test. We clearly observed evidence of publication bias in the funnel plot, as studies were oddly distributed across the effect line (Figure 4). However, we did not observe

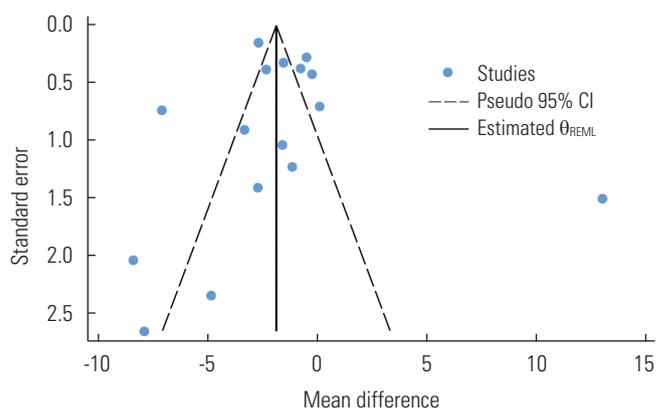


Figure 4. Funnel plot showing any evidence of publication bias. CI, confidence interval; REML, restricted maximum-likelihood estimator.

statistical evidence for the small-study effect ($p = 0.651$). The trim-and-fill analysis did not show any significant change in the effect estimate ($p = 0.358$).

DISCUSSION

The present systematic review and meta-analysis included 16 studies with a total of 5377 COVID-19-positive patients to

study the association of lower levels of vitamin D and comorbidities with the risk of COVID-19 infection. Through an analysis of the available data, we found a significant association of vitamin D deficiency with the risk of COVID-19 infection [19, 26,53-57]. A possible cause for this observation may be related to vitamin D's role in a variety of the body's immune responses. Vitamin D, also known as calcitriol, is the active form of vitamin D. Upon interacting with the vitamin D receptor (VDR), which is found on immune cells (B, T, and antigen-presenting cells) and pulmonary epithelial cells, the calcitriol-VDR complex induces transcriptional expression of antimicrobial peptides, such as cathelicidins and defensins. Cathelicidins disrupt bacterial cell membranes, as well as enveloped viruses, such as SARS-CoV-2, while defensins promote chemotaxis of inflammatory cells via increased capillary permeability. Despite the fact that vitamin D promotes the expression of several inflammatory cytokines through T-cell inactivation and interferon activation, it also inhibits the pro-inflammatory markers interleukin-6 and tumor necrosis factor- α , which are 2 key cytokines involved in the development of the "cytokine storm" that precedes acute respiratory stress disorder. Given this molecular understanding, it is reasonable to believe that individuals with vitamin D deficiency are at a high risk of developing more severe COVID-19 symptoms and/or a worse prognosis [58].

For infectious diseases caused by viruses, there are abundant and diverse ways in which sex can impact differential susceptibility between males and females. Although many studies have addressed the sex discrepancy in COVID-19, very few reports have analyzed the underlying cause of this disparity [59,60]. Several studies have analyzed levels of vitamin D in COVID-19 patients according to sex [61-63]; however, the reported findings are contradictory. We performed a subgroup analysis based on sex and observed a stronger tendency for female COVID-19 patients to have lower levels of vitamin D than COVID-19-negative patients (MD, -2.18; 95% CI, -4.08 to -0.28; $p=0.024$) than was the case for male patients (MD, -1.74; 95% CI, -3.79 to 0.31; $p=0.096$). Male patients showed higher odds of having low vitamin D levels (OR, 2.09; 95% CI, 1.38 to 3.17; $p<0.001$) than female patients (OR, 1.17; 95% CI, 0.74 to 1.86; $p=0.477$).

Vitamin D deficiency has been shown to be a risk factor for COVID-19, especially for severe/critical cases [64]. Various retrospective observational studies have demonstrated an association of vitamin D deficiency with COVID-19 risk [19,26,65] and have suggested the usefulness of vitamin D supplementa-

tion to reduce the risk of infection [18,20,26,64]. Elderly individuals and people with comorbidities are more susceptible to severe COVID-19 infection and may demonstrate worse morbidity outcomes [66-68]. Since the risk of symptomatic upper respiratory tract infection has been suggested to be associated with low vitamin D levels, its concentrations are expected to be quite low in COVID-19-positive patients [20]. Guan et al. [69] reported comorbidities and their impacts on 1590 COVID-19 patients, and indicated that 399 (25.1%) patients reported at least 1 comorbidity, including hypertension, cardiovascular disease, cerebrovascular disease, diabetes, hepatitis B infection, chronic obstructive pulmonary disease, chronic kidney disease, malignancy, and immunodeficiency. Subsequently, Wang et al. [70] reported findings from 138 cases of COVID-19; the results suggested that comorbidities may be risk factors for adverse outcomes. In our study, we found that obesity and socioeconomic status were major risk factors for COVID-19 infection, which may be supported by the findings of meta-analyses performed to investigate the association of obesity and socioeconomic status in COVID-19 patients [71,72]. Assessing the prevalence of chronic diseases forms the basis for mitigating complications in patients with COVID-19. However, these efforts were hampered by the limited number of cases in the earliest stages of the pandemic [73].

Compared with previous meta-analyses [53-57], our study had a larger sample size, making the results more credible. Furthermore, an analysis of vitamin D levels by sex was not performed in previous studies. The evidence presented in this review shows promise for the use of vitamin D supplementation to reduce the risk and severity of COVID-19 infection.

To our knowledge, this meta-analysis included the highest number of COVID-19 patients not treated with any vitamin D supplementation. The present meta-analysis demonstrated a significant association between COVID-19 positivity and vitamin D levels in case-control studies, while no association was found in cohort studies. The reason for the difference in the results may be due to differences in the study design; although observational studies are more prone to bias, it was still seen that the case-control study design showed different results from those of cohort studies. This meta-analysis of observational studies provides a general idea of the association between vitamin D levels and COVID-19. Hence, further randomized studies are recommended to be conducted to assess the effect of vitamin D on COVID-19. The present study has some limitations: (1) there are discrepancies in the number and sample size of the

included studies, leading to some instances of large variance in effect size estimates; and (2) significant heterogeneity was found, and we only used random-effects models to address heterogeneity, which may have affected the strength and extrapolation of conclusions; (3) publication bias may have affected our results because negative studies were less likely to be published; and (4) although we conducted an extensive search, we may have inadvertently missed some relevant studies.

CONCLUSION

Low serum vitamin D levels were significantly associated with a high risk of acquiring COVID-19 infection; however, the results varied by study design. This relationship was more prominent in female patients than in male patients. Limited evidence was found regarding relationships between reported comorbidities and COVID-19 infection; therefore, large population-based studies are recommended to establish any association. Vitamin D-deficient individuals should be provided special attention. Vitamin D levels can be monitored and supplementation of vitamin D could be considered in patients to improve their recovery if they contract COVID-19.

SUPPLEMENTAL MATERIALS

Supplemental material is available at <https://doi.org/10.3961/jpmp.21.640>.

CONFLICT OF INTEREST

The authors have no conflicts of interest associated with the material presented in this paper.

FUNDING

None.

ACKNOWLEDGEMENTS

Ram Bajpai is affiliated with the National Institute for Health and Care Research (NIHR) Applied Research Collaboration (ARC) West Midlands. The views expressed are those of the author(s) and not necessarily those of the NIHR or the Department of Health and Social Care.

AUTHOR CONTRIBUTIONS

Conceptualization: Mishra P, Parveen R. Data curation: Mishra P, Parveen R. Formal analysis: Mishra P, Bajpai R. Funding acquisition: None. Methodology: Mishra P, Bajpai R. Project administration: Agarwal N, Mishra P. Visualization: Mishra P, Parveen R. Writing – original draft: Mishra P, Parveen R, Agarwal N, Bajpai R. Writing – review & editing: Bajpai R, Agarwal N.

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